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# Do ecological differences between taxonomic groups influence the relationship between species' distributions and climate? A global meta-analysis using species distribution models

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Understanding whether and how ecological traits affect species' geographic distributions is a fundamental issue that bridges ecology and biogeography. While climate is thought to be the major determinant of species' distributions, there is considerable variation in the strength of species' climate–distribution relationships. One potential explanation is that species with relatively low dispersal ability cannot reach all geographic areas where climatic conditions are suitable. We tested the hypothesis that species from different taxonomic groups varied in their climate–distribution relationships because of differences in life history strategies, in particular dispersal ability. We conducted a meta-analysis by combining the discrimination ability (AUC values) from 4317 species distribution models (SDMs) using fit as an indication of the strength of the species' climate–distribution relationship. We found significant differences in the strength of species' climate–distribution relationships across taxonomic groups, however we did not find support for the dispersal hypothesis. Our results suggest that relevant ecological trait variation among broad taxonomic groups may be related to differences in species' climate–distribution relationships, however which ecological traits are important remains unclear.

Understanding whether and how ecological traits affect species' geographic distributions is a fundamental issue that bridges ecology and biogeography (Brown 1995, Wiens 2011). This issue has become even more relevant as ecologists and biogeographers struggle to understand the variation in species' responses to climatic change. For example, recent studies have examined the relationship between species' ecological traits, such as dispersal ability and ecological generalization, and changes in their distributions and phenology with recent climatic changes (Angert et al. 2011, Diamond et al. 2011). Identifying characteristics of organisms that determine their sensitivity to environmental change is crucial to ecological forecasting and conservation planning.

Central to this work is the theory of the niche: the set of abiotic and biotic conditions within which a species can persist (Hutchinson 1957). A species' distribution is limited to geographic areas where all these conditions meet the species' niche requirements. At broad spatial scales, climate has long been considered the most important factor in determining species' distribution limits (Merriam 1894, Good 1931, Gaston 2003). However, there seems to be considerable variation in the degree to which species' distributions are predicted by climate. There are three potential reasons for this variation. First, other abiotic or biotic factors may

prevent a species from persisting even where the climate is suitable (Luoto et al. 2007). Alternatively, regions of suitable climate may be separated by areas that are not suitable which the species does not have sufficient dispersal ability to cross (Blach-Overgaard et al. 2010, Graham et al. 2010). Finally, if the species is relatively new and/or the climate has only recently become suitable, the species may not have had enough time to reach all suitable areas (Paul et al. 2009, Blach-Overgaard et al. 2010).

Dispersal ability is thought by some to determine how closely a species' current distribution matches the geographic distribution where all abiotic and biotic conditions meet its niche requirements. Species that produce many propagules that travel long distances are more likely to be able to cross any unsuitable habitat, and thus should be more likely to be found everywhere the climate is suitable. Therefore, dispersal ability may determine the strength of the species' climate–distribution relationship. Indeed, some studies have found evidence that dispersal ability can strongly affect species' distributions (Thuiller et al. 2004, Poyry et al. 2008). However, others suggest that the dispersal of individuals happens over such small time scales relative to the formation of species' geographic distributions that it has little importance (Lester et al. 2007).

Many have hypothesized that species in different taxonomic groups should vary in their climate–distribution relationships because of their different life history strategies, in particular dispersal ability (Araújo and Pearson 2005, Wisz et al. 2008). The fit of species distribution models (SDMs) has often been used to test this hypothesis (Araújo and Pearson 2005, Tsoar et al. 2007). SDMs use various statistical techniques to describe the relationship between observed environmental variables, such as mean annual temperature, and the recorded spatial occurrence (presence/absence) of a species (Guisan and Zimmermann 2000). The ability of an SDM based only on climatic factors to predict the presence or absence of a species can be considered an indication of the strength of the species' climate–distribution relationship: the greater the success of a SDM at predicting the species' presence/absence in a given location, the stronger the correlation between climatic variables and the presence/absence of the species. Some studies have found species' climate–distribution relationship differences between taxonomic groups (Araújo and Pearson 2005, Tsoar et al. 2007), whereas others have not (Pearce and Ferrier 2000, Wisz et al. 2008). It is unclear whether these varying results are due to the different geographic regions, groupings of species, or modeling techniques of each study. Despite the availability of SDMs for thousands of species, a comprehensive comparison of the fit of SDMs between different taxonomic groups has not been made.

Here, we tested the hypothesis that taxonomic groups varied in the strength of their species' climate–distribution relationships. We predicted that taxonomic groups with lower dispersal ability would have weaker species' climate–distribution relationships. We used a meta-analysis approach and combined the discrimination ability metrics that were reported from 4317 SDMs in twenty studies using only climatic variables to determine whether species varied predictably in their climate–distribution relationships based on taxonomic affinities. We also compiled dispersal distances for a subset of these species to determine whether dispersal ability directly influenced the strength of species' climate–distribution relationships. To facilitate a quantitative comparison we used a standardized discrimination ability measure and accounted statistically for methodological differences among studies.

## Material and methods

### Data compilation

We conducted a literature search using Web of Science for studies (published before March 2009) that reported statistical measures of goodness-of-fit for SDMs constructed for individual species based on climatic variables only. We searched for studies using the terms 'ecological niche model' and 'climat\*', 'species distribution model' and 'climat\*', and 'climate envelope model' and 'climat\*'. Studies were excluded if: 1) one or more non-climatic variables, such as soil fertility, land use or land cover, were included in the SDM; 2) model fit was measured only qualitatively or not reported; or 3) model fit was reported only as averages across species. In cases where model fit was not reported for

all individual species modeled, we requested these data from the authors. Due to the small number of studies modeling aquatic species, we limited our analysis to terrestrial species.

We needed a metric of model fit that was comparable across studies. We found AUC (area under a receiver operating characteristic curve) to be the most common metric (other metrics included: Cohen's kappa, sensitivity, specificity, range filling rates), therefore our analysis was limited to studies that reported AUC. AUC measures the ability of a SDM to discriminate sites where a species is present from sites where it is absent, rather than goodness-of-fit per se. It considers the relationship between false-positives and true-positives and ranges from zero to one, where perfect discrimination gives a value of one (Fielding and Bell 1997). Hereafter, we use the term SDM 'fit' to indicate 'discrimination ability' as measured by AUC. When studies reported AUC for both training and test data, test AUC values were used. Although this metric has been criticised (Lobo et al. 2008), it was the only measure in common across most of the studies.

Some species' distributions were modeled several times, either by the same study (using multiple modeling techniques ( $n = 9$ ) or resolutions ( $n = 1$ )) or by several studies (most such species were modeled by only two studies). In all cases, we randomly selected one SDM per species and used the associated AUC value and methodology. This produced a dataset of 4317 species and their SDMs from twenty studies (Supplementary material Appendix 1–3). These studies modeled species in Europe (10 studies, 2301 spp.), North America (2 studies, 67 spp.), South America (2 studies, 32 spp.) and Africa (6 studies, 1917 spp.). We classified each species into one of five broad taxonomic groups: mammals (483 spp.), butterflies (116 spp.), herptiles (reptiles and amphibians; 114 spp.), birds (2099 spp.), and plants (1505 spp.).

SDM fit can be affected by the type of model used (Elith et al. 2006), the number of climatic variables used (Pearce and Ferrier 2000), the resolution or grain size used (Guisan et al. 2007), the total extent over which the species' range was modeled (Luoto et al. 2005), and latitude (Brown et al. 1996, Luoto et al. 2005). Therefore, for each SDM we noted the modeling technique, number of distinct climatic variables used in the model, resolution ( $\text{km}^2$ ), total spatial extent ( $\text{km}^2$ ) and average absolute latitude and then included these as covariates in our statistical analysis.

Another factor which may lead to differences in SDM fit between species is prevalence (McPherson et al. 2004, Santika 2011), the number of grid cells from which a species is recorded as present expressed as a proportion of the total number of grid cells from which data are available. We were able to obtain prevalence values for almost all of the SDMs ( $n = 4089$ ), allowing us to explore any effects of prevalence on SDM fit.

Finally, we scanned the literature to find dispersal distances for as many of our species as possible to assess whether there were significant differences in measured dispersal ability among our taxonomic groups. True dispersal distances are very difficult to measure due to phenomena such as very rare long-distance dispersal events. Therefore, we used the directly measured ability of an organism or its propagules to move (i.e. its mobility) as an estimate of a species' dispersal

distance. We considered both maximum and mean measured dispersal distances but excluded migratory distances to standardize measures of dispersal distances across taxonomic groups. Where more than one distance was reported per species or study we used the mean of mean distances, and the maximum of maximum distances. We found mean dispersal distances for 241 species for which we also had AUC values (birds = 103, butterflies = 22, mammals = 22, plants = 94). For maximum dispersal distance, we found 105 species that also had AUC values (birds = 27, butterflies = 18, herptiles = 12, mammals = 30, plants = 18). For further details, see Supplementary material Appendix 4, 5.

## Statistical analysis

There were two parts to the analysis. The first was to determine whether there were any significant differences in SDM fit between taxonomic groups and whether those differences were robust to potential confounding factors (covariates). The second was to explore the relationship between SDM fit, taxonomic group and the other covariates. We used generalized linear mixed-effects models (GLMM, `glmmadmb` function in the 'glmmADMB' package (Skaug et al. 2012) in R (R Development Core Team)) with a Beta error distribution with AUC as our response variable and 'study' as a random factor. AUC values of exactly one, which are not allowed with the beta distribution, were converted to 0.99 instead (eight significant digits were used to ensure a unique value and to match the maximum precision of the data,  $n = 117$ ). To allow for model estimation, we collapsed the six rarest modeling types into one category to reduce the number of types (from 18 to 12; these six techniques were used for only 0.35% of all SDMs). We took the logarithm of spatial extent to improve normality (except

in the collinearity test), but all other covariates were used without transformation. Taxonomic group and model type were categorical, and all other covariates were continuous.

## Relationship between discrimination ability and taxonomic group

To test whether taxonomic group explained significantly more deviance in AUC than expected at random, we compared a model with only an intercept to a model with only taxonomic group. We then tested whether differences in discrimination ability across taxonomic groups explained significant additional deviance after accounting for the combined effect of the differences in the methodological approach of studies (i.e. the covariates: model type, resolution, number of climatic variables, spatial extent and latitude). For all model comparisons, we used a likelihood ratio test. We also calculated AIC for all models to evaluate the relative effects of individual covariates.

We first inspected bivariate plots of all continuous covariates before constructing pairwise correlations to identify potential problems with multi-collinearity among covariates (Supplementary material Appendix 6). Latitude was highly correlated with spatial extent and resolution (Spearman's  $r = -0.903$ ,  $-0.589$  respectively,  $n = 4317$ , Supplementary material Appendix 6) and explained less deviance in AUC than spatial extent or resolution (Table 1), therefore the 'full model' included taxonomic group, model type, spatial extent resolution and number of climatic variables. We considered the effect of 'study' by including it as a random factor and by testing the influence of individual studies that contributed more than half of the total number of species in one taxonomic group ('large studies') by comparing results obtained with and without each of these studies (Araújo et al. 2005, Luoto et al. 2005, Huntley et al. 2006, Supplementary material Appendix 3).

Table 1. Analysis of deviance table for the relationship between discrimination ability, covariates and taxonomic group. Presented are the differences in degrees of freedom, AIC and deviance between full and reduced models as well as the associated p value. Models are compared for all species ( $n = 4317$ ) and for the subset of species with prevalence values ( $n = 4089$ ). Depending on the model comparison and term of interest, the full model includes all other covariates (number of variables,  $\log(\text{spatial extent})$ , model type, resolution and taxonomic group).

| Model for comparison | Data  | Model terms                     | Difference in DF | $\Delta\text{AIC}$ | $\Delta\text{Deviance}$ | p       |
|----------------------|---|---------------------------------|------------------|--------------------|-------------------------|---------|
| Just intercept       | All species                                 | Intercept                       |                  |                    |                         |         |
|                      |   | + taxonomic group               | 4                | 38.98              | 46.98                   | <0.0001 |
|                      |   | + model type                    | 10               | 100.58             | 120.58                  | <0.0001 |
|                      |   | + $\log(\text{spatial extent})$ | 1                | 2.58               | 4.58                    | 0.03235 |
|                      |   | + resolution                    | 1                | 1.38               | 3.38                    | 0.0660  |
|                      |   | + number of climatic variables  | 1                | 1.20               | 0.8                     | 0.3711  |
| Subset               | Subset                                      | + latitude                      | 1                | 0.58               | 1.42                    | 0.2334  |
|                      |   | Intercept                       |                  |                    |                         |         |
| Full model           | All species                                 | + prevalence                    | 1                | 335.36             | 337.36                  | <0.0001 |
|                      |   | Full model                      |                  |                    |                         |         |
|                      |   | + taxonomic group               | 4                | 38.64              | 46.64                   | <0.0001 |
|                      |   | + model type                    | 10               | 101.52             | 120.14                  | <0.0001 |
|                      |   | + $\log(\text{spatial extent})$ | 1                | 1.12               | 3.12                    | 0.0773  |
|                      |   | + resolution                    | 1                | -1.38              | 0.62                    | 0.431   |
|                      | + number of climatic variables <sup>a</sup> | NA                              | NA               | NA                 | NA                      |         |
| Subset               | Subset                                      | Full model                      |                  |                    |                         |         |
|                      |   | + prevalence <sup>†</sup>       | 1                | 445.62             | 447.62                  | <0.0001 |

<sup>a</sup>No solution was found.

<sup>†</sup>A model solution could only be found if number of climatic variables was not included.

### **Relationship between SDM fit, covariates and taxonomic group**

We tested whether individual covariates (including prevalence) explained significantly more deviance in AUC than under random expectation and after accounting for all other covariates (including taxonomic group) by comparing each model to a reduced one. Finally, to test whether there were significant differences in dispersal distance (both mean and maximum) across taxonomic groups, we used a Kruskal–Wallis rank sum test. We then tested whether dispersal distance explained significantly more deviance than expected by chance in AUC by comparing a model with and without dispersal distance. Dispersal distance was log-transformed to improve normality. Lastly, to test for the possibility that an interaction between dispersal distance and taxonomic group explained deviance in AUC, we compared a model with and without this two-way interaction.

All statistical analyses were performed using R 2.14.1 (R Development Core Team).

## **Results**

### **Relationship between discrimination ability and taxonomic group**

Mean AUC across all species was 0.941 ( $\pm 0.00104$  SE,  $n = 4317$ ). Birds had the highest mean AUC ( $0.954 \pm 0.00145$  SE,  $n = 2099$ ) and butterflies had the lowest mean AUC ( $0.856 \pm 0.0114$  SE,  $n = 116$ ; Fig. 1a). However, the ranking and pair-wise comparison of taxonomic groups changed depending on which ‘large study’ was removed (Fig. 1).

Taxonomic group explained significant deviance in AUC ( $LRT_{7,3} = 46.98$ ,  $p < 0.0001$ ; Table 1), even after accounting for all covariates ( $LRT_{20,16} = 46.64$ ,  $p < 0.0001$ ; Table 1). The effect of taxonomic group was also robust to the exclusion of each of the ‘large studies’ (Supplementary material Appendix 7).

### **Relationship between discrimination ability, covariates and taxonomic group**

SDM model type explained significant deviance in AUC ( $LRT_{3,13} = 120.58$ ,  $p < 0.0001$ ; Table 1), even after accounting for all the other covariates ( $LRT_{20,10} = 120.14$ ,  $p < 0.0001$ ; Table 1). For the subset of species for which we had prevalence data, prevalence also explained significant deviance in AUC after accounting for all covariates (including taxonomic group;  $LRT_{12,11} = 447.62$ ,  $p < 0.0001$ ; Table 1). SDMs with greater prevalence had lower AUC (Spearman’s  $r = -0.4937$ ).

In our subset of species with dispersal distances, mean dispersal distance was greatest for mammals (175 km) while birds had the greatest maximum dispersal distance (1305 km; Fig. 2). Butterflies had the shortest mean and maximum dispersal distance (0.441 and 2.25 km, respectively; Fig. 2). The ranking of groups closely matched the ranking of groups of the entire dataset in terms of AUC for both dispersal measures (Fig. 1a, 2). There was also a significant difference between taxonomic groups in dispersal

distance (mean:  $DF = 3$ ,  $\chi^2 = 181.006$ ,  $p < 0.0001$ ; max:  $DF = 4$ ,  $\chi^2 = 291.557$ ,  $p < 0.0001$ ). Taxonomic group explained significant deviance in AUC (mean:  $LRT_{6,3} = 10.386$ ,  $p = 0.01555$ ; max:  $LRT_{7,3} = 13.022$ ,  $p = 0.01117$ ). However, dispersal distance did not explain significant deviance in AUC (mean:  $LRT_{4,3} = 2.068$ ,  $p = 0.1504$ ; max:  $LRT_{4,3} = 0.144$ ,  $p = 0.7043$ ). There was no significant interaction between taxonomic group and dispersal distance (mean:  $LRT_{10,7} = 4.508$ ,  $p = 0.2116$ ; max:  $LRT_{12,8} = 4.506$ ,  $p = 0.3418$ ).

## **Discussion**

We found support for taxonomic differences in SDM fit suggesting a role for ecological traits in affecting species’ geographic distributions at broad scales. However, prevalence and methodological issues, such as model type, also influenced SDM fit. Indeed, both factors have been shown previously to influence SDM fit (Elith et al. 2006, Santika 2011). We also found that ‘large studies’ influenced the relationship among taxonomic groups and AUC, for example the taxonomic group with the highest mean AUC varied with the subset of species considered (Fig. 1). Therefore, species’ taxonomic affinities, prevalence and methodological issues, such as the model type, are all important in influencing species’ climate–distribution relationships as measured by SDMs.

There are a number of potential explanations for the difference in the strength of species’ climate–distribution relationships between taxonomic groups. First, taxonomic differences may reflect differences in dispersal ability among groups. Certainly, we found differences in measured dispersal distances between broad taxonomic groups that were consistent with the dispersal hypothesis (Fig. 1a, 2). However, there were inconsistencies in the ranking and pair-wise comparisons of taxonomic groups in SDM fit depending on the subset of species considered (Fig. 1). Moreover, there was no significant relationship between AUC and dispersal distance. Therefore, our results indicate that greater dispersal ability, at least in terms of measurable differences in mobility, may not result in stronger overall species’ climate–distribution relationships at broad scales. However, dispersal distance is inherently difficult to measure and our estimate of dispersal ability may not have been the most appropriate for all species. For example, we did not take into account migratory or rare long-distance dispersal events. Consequently, we may have underestimated the role of dispersal ability for certain species.

Alternatively, dispersal may not be an important trait in determining species’ climate–distribution relationships. The majority of species had low prevalence (77% species had  $< 0.1$  prevalence) and species with lower prevalence were more likely to have higher AUC values. If these low prevalence species are mainly specialists (i.e. restricted range endemics) that are adapted to uncommon climatic conditions found in small, contiguous areas, they could have strong climate–distribution relationships regardless of dispersal ability.

Third, other life history traits, for example, body size, generation time or diet breadth, may influence the strength

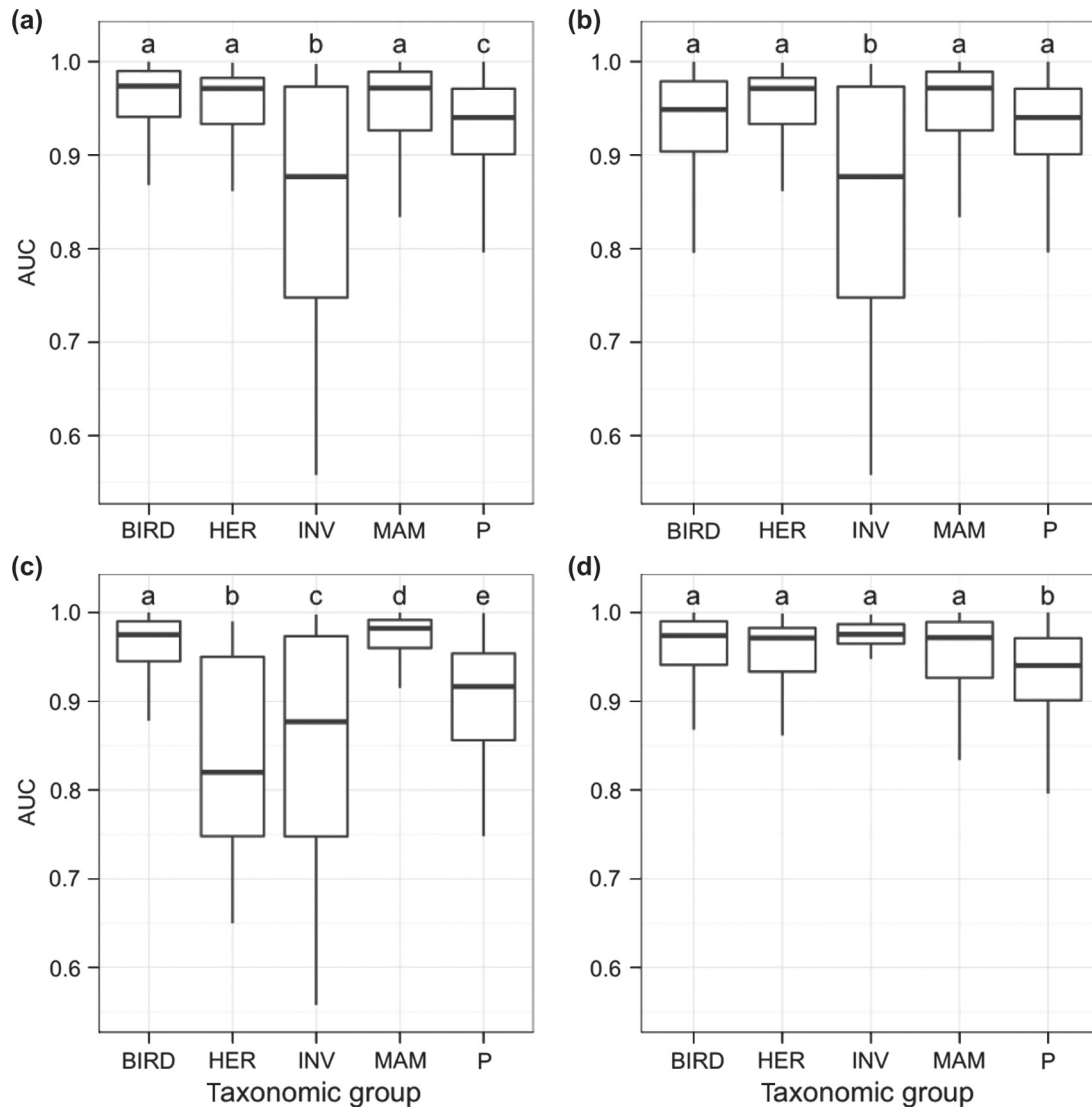


Figure 1. Taxonomic differences in discrimination ability (AUC) across all studies (based on 4317 species from twenty published studies (number of species: birds  $n = 2099$ ; herptiles  $n = 114$ ; butterflies  $n = 116$ ; mammals  $n = 483$ ; plants  $n = 1505$ )) (a), without Huntley et al. 2006 (based on 2860 species from nineteen published studies (number of species: birds  $n = 642$ ; herptiles  $n = 114$ ; butterflies  $n = 116$ ; mammals  $n = 483$ ; plants  $n = 1505$ )) (b), without Araújo et al. 2005 (based on 2539 species from nineteen published studies (number of species: birds  $n = 1942$ ; herptiles  $n = 11$ ; butterflies  $n = 116$ ; mammals  $n = 331$ ; plants  $n = 139$ )) (c), and without Luoto et al. 2005 (based on 4238 species from nineteen published studies (number of species: birds  $n = 2099$ ; herptiles  $n = 114$ ; butterflies  $n = 37$ ; mammals  $n = 483$ ; plants  $n = 1505$ )) (d), taxonomic groups represented are: 'BIRD' = birds, 'HER' = herptiles, 'INV' = butterflies, 'MAM' = mammals, 'P' = plants. Taxonomic groups with different letters above them are significantly different according to pair-wise comparisons. Outliers were removed to improve visual contrasts between taxonomic groups.

of species' climate–distribution relationships between taxonomic groups. However, determining their relative importance may be difficult across the broad taxonomic groups considered. Lower-order taxonomic groups, or functional groups of species within or across taxonomic groups, might be more effective in dividing species according to relevant traits. Nevertheless, while some recent studies dividing species into finer taxonomic or functional divisions have found significant differences in species' climate–distribution relationships (Syphard and Franklin 2010), others have not (Huntley et al. 2004).

On the other hand, taxonomic differences in SDM fit may be a function of the sample unbalance (across studies

and taxonomic groups; Supplementary material Appendix 2) and the high average discrimination ability. Both of these factors could reflect issues related to fitting, testing and publishing SDMs. SDMs have been criticized for not using independent data to test their models (Hampe 2004, Segurado et al. 2006). Without independent test occurrence points, well-fitting models could reflect spatial autocorrelation between training and testing points rather than relationships between species' presence/absence and climatic variables. Moreover, SDMs may be overfitted by fitting complex response curves and re-fitting models until a high AUC is achieved (Araújo et al. 2005, Guisan and Thuiller 2005). We also suggest that there could be a

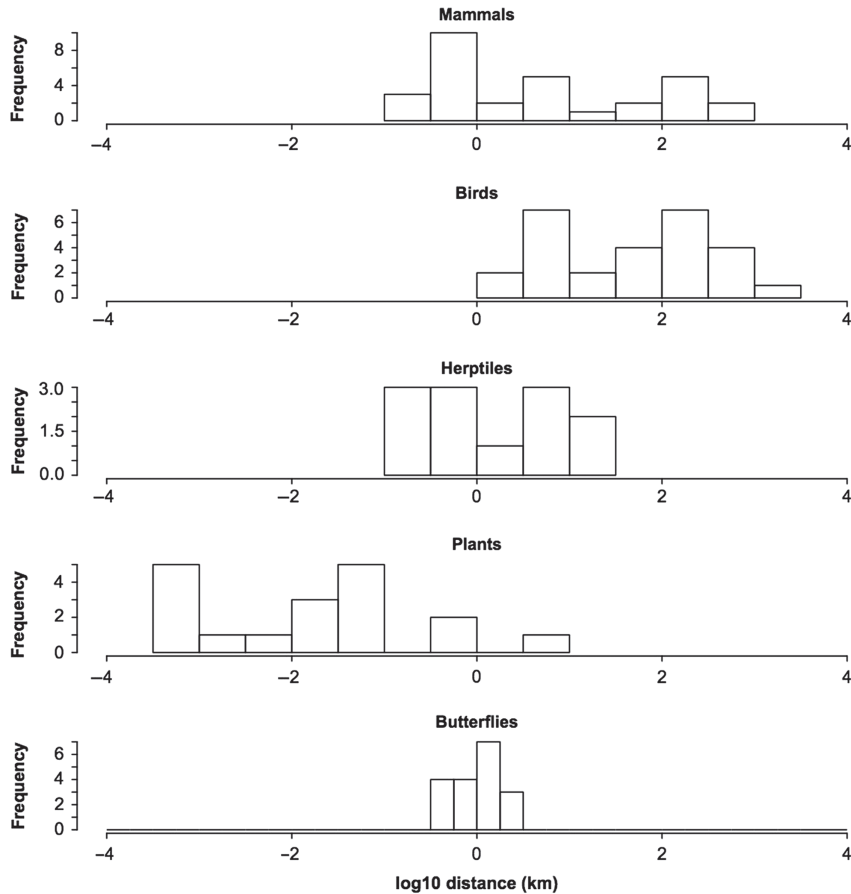


Figure 2. Taxonomic differences in log (base 10) maximum dispersal distances (km) for 105 species (birds = 27, butterflies = 18, herptiles = 12, mammals = 30, plants = 18).

‘file-drawer’ problem, whereby species that do not achieve a high enough AUC value based on the literature standard (Swets 1988) are not published. In particular, when the objective of fitting the SDM is to predict species’ potential distribution shifts under various climate change scenarios, authors (rightly) do not use SDMs with very low discrimination ability. For example, of the 453 species that Huntley et al. (2008) modeled, 13 native species that did not yield ‘useful’ models (sensu Swets 1988) were excluded from the synthesis. Taken together, these issues could inflate AUC values and reduce overall variation, making it difficult to detect the true relationship between taxonomic groups. While we acknowledge these limitations of SDMs, to our knowledge, there are no other comparable published metrics to evaluate individual species’ climate–distribution relationships at such large scales. Moreover, SDMs are still being used to better understand the relationship between species’ distributions and climate (Blach-Overgaard et al. 2010, Graham et al. 2010).

Lastly, because SDMs are fitted to species’ current distributions they reflect both direct and indirect influences of climate on those distributions. Non-climatic factors that limit a species to certain broad areas (such as biotic interactions or other abiotic factors) are generally modulated by climatic conditions. For example, since its introduction to Hawaii, avian malaria now restricts native bird species to higher elevations, where temperature halts development

of the malaria pathogen inside its mosquito vector (van Riper et al. 1986). Differences among taxonomic groups in the ability of climate to directly limit species’ distributions thus cannot be revealed by our data, given that the SDMs we used cannot differentiate direct from indirect climatic effects. However, we have no a priori reason to expect cases where climate acts principally indirectly to occur more frequently in one taxonomic group than another. In addition, even if a species’ distribution is indirectly limited by climate due to the climatic tolerances of a competitor, predator, or disease, at broad scales, climate is still the ultimate determinant of the species’ distribution.

There are a number of steps to be taken in the future to clarify how ecological traits influence species’ climate–distribution relationships. Firstly, more SDMs are needed for some taxonomic groups, particularly invertebrates and herptiles. Secondly, we should strive to eliminate issues related to species distribution modeling by using spatially/temporally independent training and test datasets where possible (Beerling et al. 1995, Randin et al. 2006). Third, analyzing SDM prediction errors might help to shed light on the mechanism driving the variation in species’ climate–distribution relationships, especially in cases of poor fit (Hanspach et al. 2011). For example, SDMs with more false negatives overall than false positives could suggest that source-sink dynamics are important: even where conditions are not favourable, individuals may still persist owing to a

rescue effect, or temporal variation in conditions (Pulliam 2000, Gaston 2003). Alternatively, models with greater rates of false positives might suggest that dispersal limitation or interspecific interactions, such as competition, are limiting a species' distribution (Pulliam 2000, Graham et al. 2010). Finally, exploring spatial variation in model behaviour, for example testing model performance in climatically heterogeneous regions or through patterns of spatial prediction errors (Hanspach et al. 2011), could also improve our understanding of model performance and thus species' climate–distribution relationships.

## Conclusion

We found a statistically significant effect of membership in broad taxonomic groups on SDM fit even after accounting for methodological issues, suggesting a role for ecological traits in determining the strength of species' climate–distribution relationships. However, the study itself, the model type used to build the SDM and species' prevalence all had significant effects on discrimination ability. Our results did not support the hypothesis that dispersal ability affects the strength of species' climate–distribution relationships. However, more work is needed to determine which ecological traits are important in determining the strength of this relationship, and at what spatial scale and taxonomic level they are manifested.

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Supplementary material (Appendix E7683 at <[www.oikosoffice.lu.se/appendix](http://www.oikosoffice.lu.se/appendix)>). Appendix 1–7.